Lesson 19: Nonparametric tests

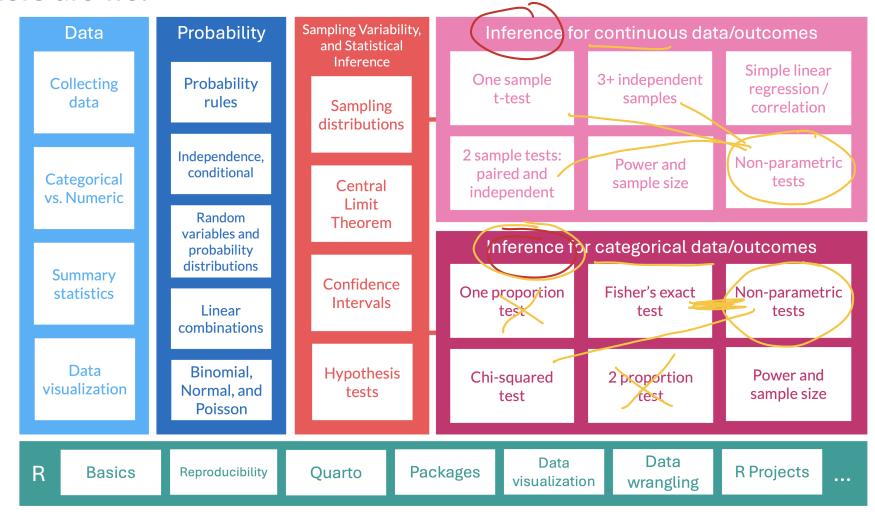
Pagona TB, Chapter 13

Meike Niederhausen and Nicky Wakim 2024-12-05

Learning Objectives

- 1. Understand the difference between and appropriate use of parametric and nonparametric tests
- 2. Use the (Wilcoxon) Signed-rank test to determine if a single sample or paired sample are symmetric around some value.
- 3. Use the Wilcoxon Rank-Sum test to compare two independent numeric samples.
- 4. Use the Fisher's Exact test to determine if two categorical variables are associated.
- 5. Use the Kruskal-Wallis test to compare two or more independent numeric samples.

Where are we?



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Background: parametric vs nonparametric

- Parametric vs. nonparametric
 - Basically saying: assuming a distribution for our data vs. not assuming a distribution for our data
- In all of our inference so far, we have assumed the population (that the data come from) has a specific distribution
 - Normal distribution, T-distribution, Chi-squared distribution, F-distribution
- Each of those distribution can be parameterized from certain population parameters
 - For example: Normal distribution is completely described (parameterized) by two parameters: μ and σ
- Our inference and analysis was all based in the assumed distribution
- Nonparametric procedures
 - Make fewer assumptions about the structure of the underlying population from which the samples were collected
 - ANOVA: equal variance Work well when distributional assumptions are in doubt.

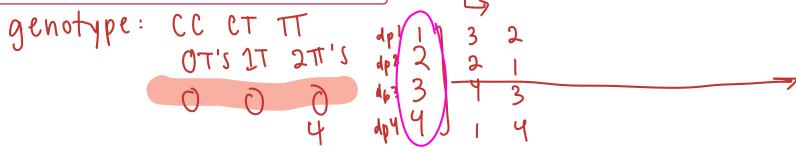
Nonparametric tests: Pros vs. cons

Pros

- Fewer assumptions
 - Can handle smaller sample sizes
 - No assumptions about the distribution of the data's population
- Tests are based on ranks Continuous
 - Therefore outliers have no greater influence than non-outliers.
 - Since tests are based on ranks we can apply these procedures to ordinal data

Cons

- Less powerful than parametric tests (due to loss of information when data are converted to ranks)
- While the test is laid out for us, it may require substantial (computer) work to develop a confidence interval
- Ties in ranks make the test harder to implement
- Some nonparametric methods can be computationally intensive, especially for large datasets or complex designs



accon 19 Slides

Parametric and nonparametric tests

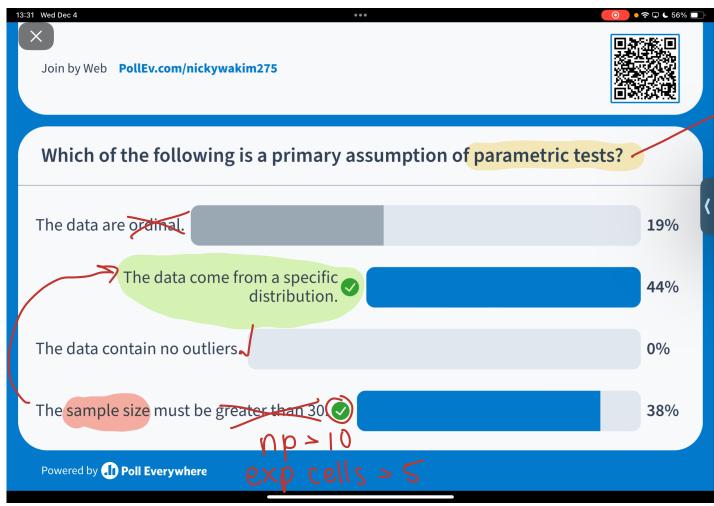
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Single sample, binary	Single proportion hypothesis test (L15)	N/A
Two independent sample, binary diff paps	Difference in proportions hypothesis test (L15)	N/A
2+ independent samples, binary	Chi-squared test (L16)	Fisher's Exact test
2+ independent samples, / numeric	ANOVA test or F-test (L17)	Kruskal-Wallis test

We still follow the general hypothesis test process

- 1. Check the assumptions
 - We will not meet the parametric assumptions!
 - There are some assumptions for the nonparametric tests
- 2. Set the level of significance α
- 3. Specify the null (H_0) and alternative (H_A) hypotheses
 - In symbols
 - In words
 - Alternative: one- or two-sided?
- 4. Calculate the test statistic and p-value
 - We will not discuss the test statistic's equation
- 5. Write a conclusion to the hypothesis test
 - Do we reject or fail to reject H_0 ?
 - Write a conclusion in the context of the problem

meet assumption > param do not meet -> noparam.

Poll Everywhere Question 1



data are..

· continuous

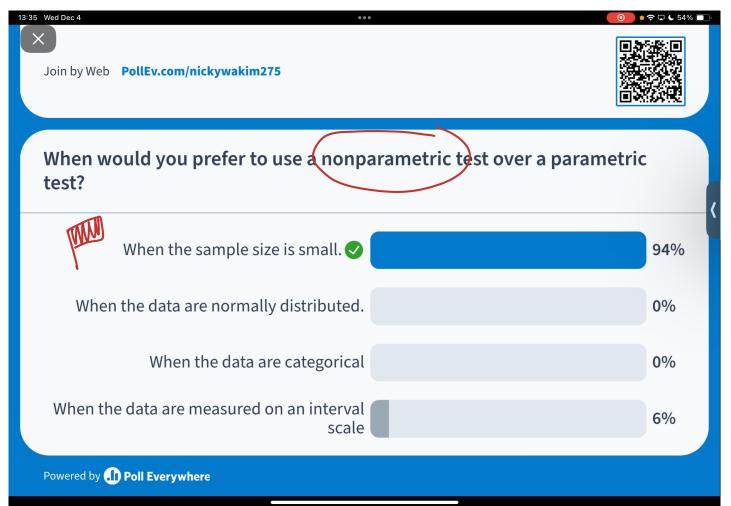
· ordinal

· binary (props)

Lesson 19 Slides

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Poll Everywhere Question 2



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(Wilcoxon) Signed-rank test

- The (Wilcoxon) signed-rank test is used for
 - Paired samples (i.e., a single set of differences)
 - One-sample comparison against a specified value Median

- If we want to see if data are symmetric (centered) around a certain value
 - For paired data, we may want to see if the data are symmetric around 0 to determine a difference
 - For one sample, we may have an idea of a median value that our data may follow

- Think back to the parametric parallel of these!
 - If we apply the body temperature example to this: We would check if the data were symmetric around 98.6

Data do NOT need to be approximately normal

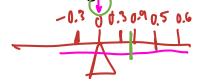
Example: Intraocular pressure of glaucoma patients

- Intraocular pressure of glaucoma patients is often reduced by treatment with adrenaline.
- A **new synthetic drug** is being considered, but it is more expensive than the **current adrenaline alternative**.
- 7 glaucoma patients were treated with both drugs:
 - one eye with adrenaline and
 - the other with the synthetic drug
- Reduction in pressure was recorded in each eye after following treatment (larger numbers indicate greater reduction)

7 < 30 × parametric

Patie	ent A	dren S	Synth	d S	Sign	
	1	3.5	3.2	-0.3)	$\left(-\right)$	1
	2	2.6	3.1	0.5	+	-
	3	3.0	3.3	0.3	+	3
	4	1.9	2.4	0.5	+	_
	5	2.9	2.9	0.0	NA	2
	6	2.4	2.8	0.4	+	4
	7	2.0	2.6	0.6	+	6

- d is the difference in reduction of pressure: Synth - Adren
- Sign is + if the difference is positive and
 if the difference is negative



(Wilcoxon) Signed-rank test: Hypotheses

General wording for hypotheses

 H_0 : population is symmetric around some value $\tilde{\mu}_0$ H_a : population is not symmetric around some value μ_0

Hypotheses test for example

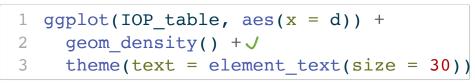
 H_0 : the population difference in reduction of intraocular pressure in treatment with adrenaline vs. new synthetic drug is symmetric around 0 H_a : the population difference in reduction of intraocular pressure in treatment with adrenaline vs. new synthetic drug is not symmetric around

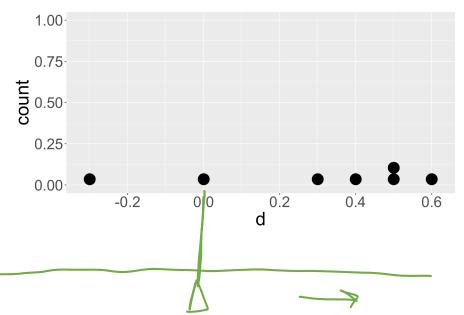
- Even if the population has a mean/median equal to $\tilde{\mu}_0$, the test may reject the null if the population isn't 50% or one side other symmetric, thus only use if the data (differences) are symmetric.
- If the population is symmetric
 - then the mean and median coincide,
 - thus often the null hypothesis is phrased in terms of the median (or median difference) being 0

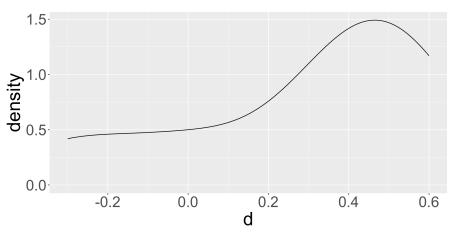
Example: Visualize the differences

Visualize the differences in reduction of pressure d: Synth - Adren

```
1 ggplot(IOP_table, aes(x = d)) +
2   geom_dotplot() +
3   theme(text = element_text(size = 30))
```







Example: Calculate signed ranks

dont get too caught up in details here.

- Rank the absolute values of the differences from smallest to largest
- For ties, take the average of the highest and lowest tied ranks
 - I.e. if ranks 3-7 are tied, then assign (3+7)/2 = 5 as the rank

• Then calculate the **signed ranks** as +/- the rank depending on whether the sign is +/-

			_	V	V	\bigvee		
	Patient	Adren	Synth	d :	Sign I	Rank S	igned_	rank
	1	3.5	3.2	-0.3	$\overline{\mathbf{C}}$	<u>+1.5</u> .		-1.5
14	2	2.6	3.1	0.5	+	4.5		4.5
J	, 3	3.0	3.3	0.3) +	1.5 •		1.5
`5	4	1.9	2.4	0.5) +	4.5		4.5
	5	2.9	2.9	0.0	NA	(NA)		NA
3	6	2.4	2.8	0.4	+	3.0		3.0
	7	2.0	2.6	0.6	+ (6.0		6.0

(Wilcoxon) Signed-rank test: Test statistic

- If the null is true:
 - lacktriangle The population is symmetric around some point ($ilde{\mu}_0=0$, typically), and
 - The overall size of the positive ranks should be about the same as the overall size of negative ranks.
- We can split the positive and negative ranks
 - \blacksquare T^+ = sum of the positive ranks
 - T^{-} sum of the negative ranks



- Thus, any of the following can be used as a test statistic and will lead to the same conclusion:
 - (T^+) what R is using)
 - $\P T^-$
 - $T^+ T^- = 0$
 - $ullet \min(T^+,|T^-|)=T_0$

Example: calculate sums of signed ranks

• Sum of the positive ranks

$$T^+ = 1.5 + 3 + 4.5 + 4.5 + 6 = 19.5$$

- Sum of the negative ranks
 - *T*⁻ = -1.5

 $ullet \min(T^+, |T^-|) = T_0 = 1.5$

Patient A	Adren S	Synth	d :	Sign I	Rank S	Signed_	rank
1		3.2			1.5		-1.5
2	2.6	3.1	0.5	+	4.5		4.5
3	3.0	3.3	0.3	+	1.5		1.5
4	1.9	2.4	0.5	+	4.5		4.5
5	2.9	2.9	0.0	NA	NA		NA
6	2.4	2.8	0.4	+	3.0		3.0
7	2.0	2.6	0.6	+	6.0		6.0

(Wilcoxon) Signed-rank test: Exact p-value (fyi)

- Exact p-value is preferable
 - This is the default method in R's wilcox.test()
 - if the samples contain less than 50 finite values
 - and there are no ties
 - R wil<mark>l automatically use nor</mark>mal approximation method if there are ties
- We will not be calculating the exact p-value "by hand." We will be using R for this.

$$p-value=2*P(allaway A) \leq t)$$

- t is the smaller of the calculated sums of the positive and negative ranks
- To calculate the exact p-value, we need the probability of each possible sum of ranks

(Wilcoxon) Signed-rank test in R: Glaucoma example

"Attempt" with exact p-value & running one sample test with differences

```
# Exact p-value
wilcox.test(x = IOP$d,
alternative = c("two.sided"), mu = 0,
exact = TRUE, correct = TRUE)

# Exact p-value
difference
Ho: Symm around

| Symm around | Open | Op
```

Wilcoxon signed rank test with continuity correction

```
data: IOP$d
V = 19.5, p-value = 0.07314
alternative hypothesis: true location is not equal to 0
```

(Wilcoxon) Signed-rank test: Conclusion

Recall the hypotheses to the (Wilcoxon) Signed-rank test:

 H_0 : the population difference in reduction of intraocular pressure in treatment with adrenaline vs. new synthetic drug is **symmetric around** $ilde{\mu}_0=0$

 H_a : the population difference in reduction of intraocular pressure in treatment with adrenaline vs. new synthetic drug is **not symmetric around** $\tilde{\mu}_0=0$

- Significance level: α = 0.05
- p-value = 0.07314
- Do we reject or fail to reject H_0 ?

Conclusion:

There is insufficient evidence the differences in reduction in intraocular pressure differs between the synthetic drug and adrenaline are symmetric about 0 (2-sided Wilcoxon signed rank test p-value = 0.07314)

(Wilcoxon) Signed-rank test with one sample

- One can use the (Wilcoxon) Signed-rank test when testing just one sample
- Note that we did this when in R: Ran the (Wilcoxon) Signed-rank test using just the differences
- For one sample, we can specify the null population median value:

 H_0 : The population median is m

 H_a : The population median is NOT m

Not-so-real example: Run (Wilcoxon) Signed-rank test for paired data with null m=0.7

```
1 wilcox.test(x = IOP$d, mu = 0.7, alternative = "two.sided")
```

Wilcoxon signed rank test with continuity correction

```
data: IOP$d V = 0, p-value = 0.02225 alternative hypothesis: true location is not equal to 0.7
```

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Wilcoxon rank-sum test

- ullet The nonparametric alternative to the two-sample t-test
 - used to analyze two samples selected from separate (independent) populations
- Also called the Mann-Whitney U test
- Unlike the signed-rank test, there is no need to assume symmetry
- Necessary condition is that the two populations being compared
 - have the same shape (both right skewed, both left skewed, or both symmetric),
 - so any noted difference is due to a shift in the median
- Since they have the same shape, when summarizing the test, we can describe the results in terms of a difference in medians.

Hypotheses:

 H_0 : the two populations have the same median

 H_a : the two populations do NOT have the same median

Example

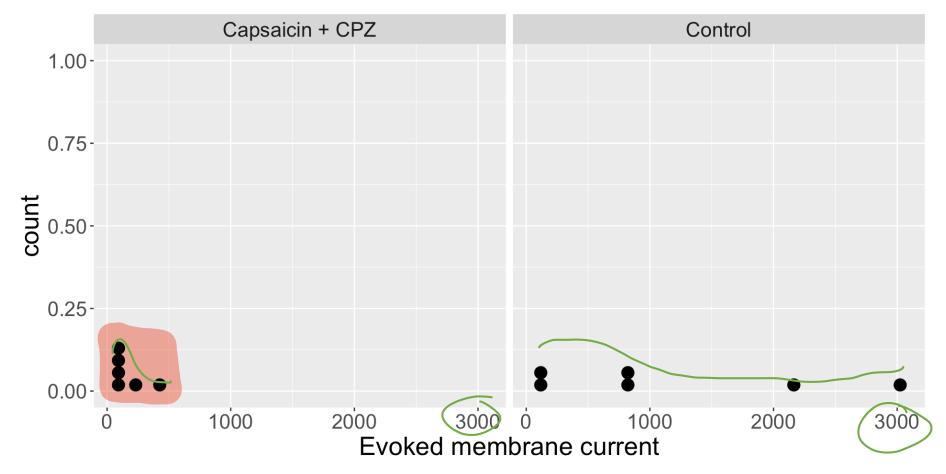
Dr. Priya Chaudhary (OHSU) examined the evoked membrane current of dental sensory neurons (in rats) under control conditions and a mixture of capsaicin plus capsazepine (CPZ). J. Dental Research 80:1518-23, 2001.

Group	variable	nı	median
Capsaici	n + CPZ Memb_current	6	112
Control	Memb_current	6	822

Rat_ID Group	Current
1 Control	3024
2 Control	2164
3 Control	864
4 Control	780
5 Control	125
6 Control	110
7 Capsaicin + CPZ	426
8 Capsaicin + CPZ	232
9 Capsaicin + CPZ	130
10 Capsaicin + CPZ	94
11 Capsaicin + CPZ	75
12 Capsaicin + CPZ	55

Example: Visualize the data

Do the independent samples have the same distribution? $\sqrt{}$



Wilcoxon rank-sum test: Calculating test statistic ${\it W}$

- 1. Combine the two samples together (keep track of which observations came from each sample).
- came from each sample). 2. Rank the full set of $N = n_1 + n_2$ observations.
 - If ties exist, assign average ranks to tied values (like signed-rank test)
- 3. Sum the ranks corresponding to those observations from the smaller sample.
 - This is a time-saving step; you could also have used the larger sample.
 - ullet Call this sum W
- 4. If n_1, n_2 are both less than 10, then use an exact test (can only be done if no ties are present)
 - Otherwise use the large-sample normal approximation.

In our example, both groups have equal n; choose either for computing W.

$W_{CPZ} = 1 + 2 + 3 + 6 + 7 + 8 = 27$	
$W_{control} = 4 + 5 + 9 + 10 + 11 + 12 = 5$	51

R	at_ID	Group	Current	Rank
	12	Capsaicin + CPZ	55	1
	11	Capsaicin + CPZ	75	2
	10	Capsaicin + CPZ	94	3
	6	Control	110	4
	5	Control	125	5
	9	Capsaicin + CPZ	130	6
	8	Capsaicin + CPZ	232	7
	7	Capsaicin + CPZ	426	8
	4	Control	780	9
	3	Control	864	10
	2	Control	2164	11
	1	Control	3024	12



Wilcoxon rank-sum test: Exact p-value approach

- If n_1, n_2 are both less than 10, then use an exact test,
 - otherwise use the large-sample normal approximation.
 - However, exact method can only be done if no ties are present
- ullet p-value is the probability of getting a rank sum W as extreme or more extreme than the observed one.
 - Multiply the 1-tail probability by 2 for the 2-tailed probability

$$p-value = 2 \cdot P(W_{CPZ} \le 27)$$

- To calculate $P(W_{CPZ} \leq 27)$,
 - we need to enumerate all possible sets ranks for the sample size,
 - calculate the sum of ranks for each set of possible ranks,
 - and then the probability for each sum (assuming each set of ranks is equally likely).

Wilcoxon rank-sum test: using R

Wilcoxon rank sum exact test

```
data: Current by Group
W = 6, p-value = 0.06494
alternative hypothesis: true location shift is not equal to 0
```

Wilcoxon rank-sum test: Conclusion

Recall the hypotheses to the Wilcoxon rank-sum test:

 H_0 : the control and treated populations have the same median

 H_a : the control and treated populations do NOT have the same median

- Significance level: α = 0.05
- p-value = 0.06494
- Do we reject or fail to reject H_0 ?

Conclusion:

insufficient evidence

There is suggestive but inconclusive evidence that the evoked membrane current of dental sensory neurons (in rats) differs between the control group and the group exposed to a mixture of capsaicin plus capsazepine (2-sided Wilcoxon rank-sum test p-value = 0.06494).

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Fisher's Exact Test

- Only necessary when expected counts in one or more cells is less than 5
- Given row and column totals fixed, computes exact probability that we observe our data or more extreme data
- Consider a general 2 x 2 table:

Group	Outco	Total	
	MANYAR	AtiveNo	
Treatment	a 12	b 43	a+b 5
Control	С	d	c+d
Total	a+c	b+d	n

• The exact probability of observing a table with cells (a, b, c, d) can be computed based on the hypergeometric distribution

$$P(a,b,c,d) = rac{(a+b)!\cdot (c+d)!\cdot (a+c)!\cdot (b+d)!}{n!\cdot a!\cdot b!\cdot c!\cdot d!}$$

Numerator is fixed and denominator changes

Some notes on the Fisher's Exact Test

- This is always a two-sided test
- There is no test statistic nor CI
- There is no continuity correction since the hypergeometric distribution is discrete

Recall our example from Lesson 4 and 16

Question: Is there an association between age group and hypertension?

• Let's pretend that we actually had the following numbers

Table: Contingency table showing hypertension status and age group.

Age Group	Hypertension	No Hypertension	Total	7
18-39 yrs	$\sqrt{1}$	11	12	
40-59 yrs	(4)	9	13	1
60+ yrs	4	$\overline{2}$	6	
Total	9	22	31	

Fisher's Exact test: Hypertension

1. Check expected cell counts threshold

We're going to pretend they are less than 5.

$$2. \alpha = 0.05$$

3. Hypothesis test:

- H_0 : There is no association between age group and hypertension
- H_1 : There is an association between age group and hypertension

4. Calculate the trestatistic and p-value for this F.E. +Cst separed tests in R

```
1 fisher.test(x = hyp_data2)
```

Fisher's Exact Test for Count Data

```
data: hyp_data2
p-value = 0.04062
alternative hypothesis: two.sided
```

5. Conclusion to the hypothesis test

We reject the null hypothesis that age group and hypertension are not associated (p=0.04062). There is sufficient evidence that age group and hypertension are associated.

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Kruskal-Wallis test: nonparametric ANOVA test

- Recall that an ANOVA tests means from 2 or more groups
- Conditions for ANOVA include
 - Sample sizes in each group group are large (each $n \geq 30$),
 - OR the data are relatively normally distributed in each group
 - Variability is "similar" in all group groups
- If these conditions are in doubt, or if the response is ordinal, then the Kruskal-Wallis test is an alternative.

```
H_0: \operatorname{pop} \operatorname{median}_1 = \operatorname{pop} \operatorname{median}_2 = \ldots = \operatorname{pop} \operatorname{median}_k vs. H_A: \operatorname{At\ least\ one\ pair\ pop\ median}_i 
eq \operatorname{pop\ median}_i 
eq \operatorname{pop\ median}_j \operatorname{for\ } i 
eq j
```

- K-W test is an extension of the (Wilcoxon) rank-sum test to more than 2 groups
 - With k=2 groups, the K-W test is the same as the rank-sum test

Ranks for the K-W test

- 1. Combine the k samples together (keep track of which observations came from each sample).
- 2. Rank the full set of $N = n_1 + \ldots + n_k$ observations.
 - If ties exist, assign average ranks to the tied values (as with the signed-rank test).
- 3. Then sum the ranks within each of the k groups
 - Label the sums of the ranks for each group as $R_1,\ldots+R_k$.

If H_0 is true, we expect the populations to have the same medians, and thus the ranks to be similar as well.

3 grp

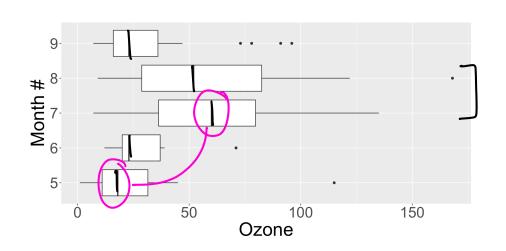
$$R_1 = \text{sum of ranks in gp1}$$
 $R_1 \approx R_2 \approx R_3$

if all have same median

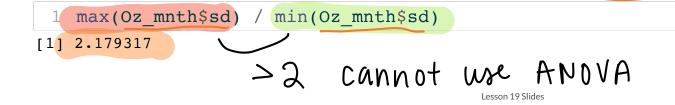
Example: Ozone levels by month

- airquality data included in base R no need to load it
- Daily air quality measurements in New York, May to September 1973.
- Question: Do ozone levels differ by month?

				lacksquare	
Month	v <u>ariabl</u> e	n) mean	median	sd
5	Ozone	26	23.615	18	22.224
6	Ozone	9	29.444	23	18.208
7	Ozone	26	59.115	60	31.636
8	Ozone	26	59.962	52	39.681
9	Ozone	29	31.448	23	24.142



• Does not look like each month has equal variance so we cannot use ANOVA



K-W test in R

outcome category gros

```
1 kruskal.test(Ozone ~ Month, data = airquality)
```

Kruskal-Wallis rank sum test

```
data: Ozone by Month
Kruskal-Wallis chi-squared = 29.267, df = 4, p-value = 6.901e-06
```

There is sufficient evidence that the median ozone levels are different in at least two months from May - September, 1973 in New York City (p < 0.001; Kruskal-Wallis test).

• (fyi) Since the K-W test is significant, follow-up with pairwise (Wilcoxon) rank-sum tests using a multiple comparison procedure to identify which months have different medians.