Lesson 3: Measurement of Association for Contingency Tables

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Lesson 3: Measurement of Association for Contingency Tables

Poll Everywhere Question 1

Make sure to remember your answer!! We'll use this on Wednesday!

Learning Objectives

1. Understand the difference between testing for association and measuring association

2. Estimate the risk difference (and its confidence interval) from a contingency table and interpret the estimate.

- 3. Estimate the risk ratio (and its confidence interval) from a contingency table and interpret the estimate.
- 4. Estimate the odds ratio (and its confidence interval) from a contingency table and interpret the estimate.

Learning Objectives

1. Understand the difference between testing for association and measuring association

- 2. Estimate the risk difference (and its confidence interval) from a contingency table and interpret the estimate.
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Review of Test of Association (1/2)

• Last week: learned some tests of association for contingency tables

- For studies with two independent samples
 - General association
 - Chi-squared test
 - Fisher's Exact test
 - Test of trends
 - Cochran-Armitage test
 - $\circ~$ Mantel-Haenszel test



Lesson 3: Measurement of Association for Contingency Tables

Test of association does not measure association

- Test of association does not provide an effective measure of association. The p-value alone is not enough
 - p-value < 0.05 suggests there is a statistically significant association between the group and outcome
 - p-value = 0.00001 vs. p-value = 0.01 does not mean the magnistude of association is different
- But it does not tell how different the risks are between the two groups
- We want to find out one or more measurements for quantifying the risks across the groups.

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Measures of Association

• When we have a 2x2 contingency table and independent samples, we have an option of three measures of association:

1. Risk difference (RD)

2. Relative risk (RR)

3. Odds ratio (OR)

- diff for each Each measures association by comparing the proportion of successes/failures from each categorical group of our explanatory variable.

Before we discuss each further...

Let's define the cells within a 2x2 contingency table:

Explanatory	Response	Total		
Variable	Success 1	Failure		
	n ₁₁	n ₁₂	n ₁	
2	n ₂₁	n ₂₂	n ₂	
Total	n ₊ (or n _s)	n ₋ (or n _F)	n	

• Then we can define risk: the proportion of "successes"

• With $\operatorname{Risk}_1 = \frac{n_{11}}{n_1}$ Ly Risk for exp group 1: $\frac{n_{11}}{n_1}$

Learning Objectives

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- 4. Estimate the odds ratio (and its confidence interval) from a contingency table and interpret the estimate.

Risk Difference (RD)

- Risk difference computes the absolute difference in risk for the two groups (from the explanatory variable)
- Point estimate:
 - Point estimate: $\widehat{RD} = \widehat{p_1} \widehat{p_1} = \frac{n_{11}}{n_1} \frac{n_{21}}{n_2}$ With range of point estimate from [-1, 1] proportion of successes in each explan. grap

Risk1

Risk2

• Approximate standard error:

$$SE_{\widehat{RD}} = \sqrt{\frac{\hat{p}_1 \cdot (1 - \hat{p}_1)}{n_1}} \pm \frac{\hat{p}_2 \cdot (1 - \hat{p}_2)}{n_2}$$
• 95% Wald confidence interval for \widehat{RD} :
Normal of approx of $\widehat{RD} \pm 1.96 \cdot SE_{\widehat{RD}}$
binomial $\widehat{RD} \pm 1.96 \cdot SE_{\widehat{RD}}$

Recall the Strong Heart Study

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The Strong Heart Study is an ongoing study of American Indians residing in 13 tribal communities in three geographic areas (AZ, OK, and SD/ND). We will look at data from this study examining the **incidence of diabetes** at a follow-up visit and **impaired glucose tolerance** (ITG) at baseline (4 years apart).

		\checkmark	0000	V -
0		Diabe	etes	
	Glucose tolerance	No	Yes	Total
	Impaired	334	198	532
	Normal	1004	128	1132
	Total	1338	326	1664



SHS Example: Risk Difference

Risk difference

Compute the point estimate and 95% confidence interval for the diabetes risk difference between impaired and normal glucose tolerance.

Needed steps:

- 1. Compute the risk difference
- 2. Compute 95% confidence interval
- 3. Interpret the estimate

	Diabe		
Glucose tolerance	No	Yes	Total
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

SHS Example: Risk Difference (1/4)

Risk difference

Compute the point estimate and 95% confidence interval for the diabetes risk difference between impaired and normal glucose tolerance.

1. Compute the risk difference

$$\widehat{RD} = \widehat{p}_1 - \widehat{p}_2 = \frac{n_{11}}{n_1} - \frac{n_{21}}{n_2} = \underbrace{\frac{198}{532}}_{532} - \underbrace{\frac{128}{1132}}_{1132} = 0.3722 - 0.1131 = \underbrace{0.2591}_{0.2591}$$

DiabetesGlucose toleranceNoYesTotalImpaired334198532Normal1004128132Total13383261664

SHS Example: Risk Difference (2/4)

Risk difference

Compute the point estimate and 95% confidence interval for the diabetes risk difference between impaired and normal glucose tolerance.

2. Compute 95% confidence interval

	Diabe		
Glucose tolerance	No	Yes	Total
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

 $egin{aligned} \widehat{RD} \pm z^*_{\left(1-rac{lpha}{2}
ight)} & imes SE_{\widehat{RD}} \ =& \widehat{RD} \pm z^*_{\left(1-rac{lpha}{2}
ight)} & imes \sqrt{rac{\hat{p}_1 \left(1-\hat{p}_1
ight)}{n_1} + rac{\hat{p}_2 (1-\hat{p}_2)}{n_2}} \ =& 0.2591 \pm 1.96 imes \sqrt{rac{0.3722 (1-0.3722)}{532} + rac{0.1131 (1-0.1131)}{1132}} \ =& (0.2141, \ 0.3041) \end{aligned}$

Risk difference

Compute the point estimate and 95% confidence interval for the diabetes risk difference between impaired and normal glucose tolerance.

1/2. Compute risk difference and 95% confidence interval

riskdifference(198, 128, 532, 1132) Mai, fmsb: Cases People at risk Risk SUCC 1982000000 532.0000000 Exposed 0.3721805 for gro] 128 0000000 1132.0000000 0.1130742 Unexposed 1664.0000000 0.1959135TOTA STotal 326.0000000 Risk difference and its significance probability (H0: The difference equals to zero) dplyr:: select() data: 198 128 532 1132 p-value < 2.2e-1695 percent confidence interval: > 95% () 0.2140779 0.3041346 sample estimates: 0.2591062 1/1

SHS Example: Risk Difference (3/4)



SHS Example: Risk Difference (4/4)

	Diabe		
Glucose tolerance	No	Yes	Total
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

Risk difference

Compute the point estimate and 95% confidence interval for the diabetes risk difference between impaired and normal glucose tolerance.

3. Interpret the estimate

The diabetes diagnosis risk difference between impaired and normal glucose tolerance is 0.2591 (95% CI: 0.2141, 0.3041). Since the 95% confidence interval contains 0, we do not have sufficient evidence that the risk of diabetes diagnosis between impaired and normal glucose tolerance is different. Beaulin 9 yr follow-up 95% CI Not contain 0: Sufficient evidence that Risk of Mat Risk of

When is the risk difference misleading?

• The same risk differences can have very different clinical meanings depending on the risk for each group

- Example: for two treatments A and B, we know the risk difference (RD) is 0.009. Is it a meaningful difference?
 - If the risk is 0.01 for Trt A and 0.001 for Trt B?
 - If the risk is 0.41 for Trt A and 0.401 for Trt B?
- Using the RD alone to summarize the difference in risks for comparing the two groups can be **misleading**
 - The ratio of risk can provide an informative descriptive measure of the "relative risk"

Learning Objectives

1. Understand the difference between testing for association and measuring association

2. Estimate the risk difference (and its confidence interval) from a contingency table and interpret the estimate.

3. Estimate the risk ratio (and its confidence interval) from a contingency table and interpret the estimate.

4. Estimate the odds ratio (and its confidence interval) from a contingency table and interpret the estimate.

Relative Risk (RR)

- Relative risk computes the ratio of each group's proportions of "success"
 - Also called risk ratio
- Point estimate:



Poll Everywhere Question 2



Log-transformation of RR

- Sampling distribution of the relative risk is highly skewed unless sample sizes are quite large \mathbb{R}
 - Log transformation results in approximately normal distribution
 - Thus, compute confidence interval using normally distributed, log-transformed RR
 - Then we convert back to the RR
- We take the log (natural log) of RR: $\ln(\widehat{RR})$ or $log(\widehat{RR})$
 - Whenever I say "log" I mean natural log (very common in statistics) Q = 2.
- Then we need to find approximate standard error for $\ln(\widehat{RR})$

$$SE_{\ln(\widehat{RR})} = \sqrt{rac{1}{n_{11}} - rac{1}{n_1} + rac{1}{n_{21}} - rac{1}{n_2}}$$

• 95% confidence interval for $\ln(\widehat{RR})$:

$$\underbrace{\ln(\widehat{RR})}_{\text{provestimal}} \pm 1.96 \times SE_{\ln(\widehat{RR})}$$

Explanatory	Response	Total		
Variable	Success	Failure		
1	n ₁₁	n ₁₂	n ₁	
2	n ₂₁	n ₂₂	n ₂	
Total	n ₊ (or n _s)	n ₋ (or n _F)	n	



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How do we get back to the RR scale?

• We computed confidence interval using normally distributed, log-transformed RR ($\ln(\widehat{RR})$):

$$\begin{pmatrix} \ln(\widehat{RR}) - 1.96 \times SE_{\ln(\widehat{RR})}, \ln(\widehat{RR}) + 1.96 \times SE_{\ln(\widehat{RR})} \\ \downarrow C \downarrow & \bigcup C \downarrow \end{pmatrix}$$

- Now we need to exponentiate the CI to get back to interpretable values
 - Take exponential of lower and upper bounds
- 95% confidence interval for RR: two ways to display equation

$$\left(\underbrace{e^{\ln(\widehat{RR}) - 1.96 \times SE_{\ln(\widehat{RR})}}_{\ln(\widehat{RR})}, e^{\ln(\widehat{RR}) + 1.96 \times SE_{\ln(\widehat{RR})}}_{\ln(\widehat{RR})} \right)$$

Relative Risk (RR)

- Can you compute the estimated RRs for the previous example?

 - If the risk for Trt A is 0.01 and Trt B is 0.001? $\widehat{RR} = 10$ Risk for trt A is 10 x If the risk for Trt A is 0.41 and Trt B is 0.401? $\widehat{RR} = 1.02$ Rore than trt B
- When $\widehat{RR} = 1$...
 - Risk is the same for the two groups
 - In other words, the group and the outcome are independent
- When computing RR it is important to identify which variable is the response variable and which is explanatory variable
 - We may say "risk for Trt A" but this translates to the risk (or probability) of outcome success for those receiving Trt A

SHS Example: Relative Risk (1/6)

Relative risk

Compute the point estimate and <u>95% confidence</u> interval for the diabetes Relative risk between impaired and normal glucose tolerance.

Needed steps:

- 1. Compute the relative risk
- 2. Find confidence interval of log RR
- 3. Convert back to RR
- 4. Interpret the estimate

	Diabe		
Glucose tolerance	No	Yes	Total
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

SHS Example: Relative Risk (2/6)

Relative risk

Compute the point estimate and 95% confidence interval for the diabetes Relative risk between impaired and normal glucose tolerance.

1. Compute the relative risk

$$\widehat{RR} = rac{\hat{p}_1}{\hat{p}_2} = rac{n_{11}/n_1}{n_{21}/n_2} = rac{198/532}{128/1132} = rac{0.3722}{0.1131} = 3.2915$$

	Diabe	Diabetes				
Glucose tolerance	No	Yes	Total			
Impaired	334	198	532			
Normal	1004	128	1132			
Total	1338	326	1664			

SHS Example: Relative Risk (3/6)

Relative risk

Compute the point estimate and 95% confidence interval for the diabetes Relative risk between impaired and normal glucose tolerance.

2. Find confidence interval of log RR

	Diabe			
Glucose tolerance	e No Yes		Total	
Impaired	334	198	532	
Normal	1004	128	1132	
Total	1338	326	1664	

 $egin{aligned} &\ln(\widehat{RR})\pm 1.96 imes SE_{\ln(\widehat{RR})} \ =&\ln(\widehat{RR})\pm z^*_{\left(1-rac{lpha}{2}
ight)} imes \sqrt{rac{1}{n_{11}}-rac{1}{n_1}+rac{1}{n_{21}}-rac{1}{n_2}} \ =&1.1913\pm 1.96 imes \sqrt{rac{1}{198}-rac{1}{532}+rac{1}{128}-rac{1}{1132}} \ =&(0.9944,\ 1.3883) \end{aligned}$

SHS Example: Relative Risk (4/6)

Relative risk

Compute the point estimate and 95% confidence interval for the diabetes Relative risk between impaired and normal glucose tolerance.

3. Convert back to RR

	Diabe		
Glucose tolerance	No Yes		Total
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

 $(\exp(0.9944), \exp(1.3883))$ =(2.703, 4.0081)

SHS Example: Relative Risk (5/6)

		Diabetes
Relative risk	Glucose tolerance	No Yes Total
Compute the point estimate and 95% confidence interval for the diabetes Relative risk between impaired and normal glucose tolerance.	Impaired Normal Total	334 198 532 1004 128 1132 1338 326 1664
1/2/3. Compute risk ratio and 95% confidence interval	glucimp	case
<pre>1 library(epitools) 2 SHS_ct = table(SHS)glucimp, SHS\$case) 3 riskratio(x = SHS_ct, rev = "rows")\$measure</pre>	norm	11
risk ratio with 95% C.I. estimate lower upper Normal 1.000000 NA NA Impaired 3.291471 2.702998 4.008061	norm L	0
Stop of RR & impair	eok	

Pause: other option in pubh package



SHS Example: Relative Risk (6/6)

				Diabe	etes	
	Relative risk		Glucose tolerance	No	Yes	Total
	$C_{\rm example}$ the neighborhood structure and 05% could be a sinter valid to the disk stars		Impaired	334	198	532
	Compute the point estimate and 95% confidence interval for the diabetes		Normal	1004	128	1132
	Relative risk between impaired and normal glucose tolerance.		Total	1338	326	1664
(3. Interpret the estimate $@$ 4 yr follow up $\widehat{RR} = 3.2$	29:	$\frac{\hat{p}_1}{\hat{p}_2}$ for \hat{p}_2 for	no	m	rd y
	The estimated risk of diabetes is 3.29 <mark>times greater</mark> for American Indians wh	o had	impaired gluc	ose to	olera	ince a
6	baseline compared to those who had normal glucose tolerance (95% CI: 2.70	, 4.01).			
	/ fraction of RR	•				
	Additional interpretation of 95% CL (not peoded): M/o are 95% confident that	t tha	nonulation) re	Jativ	- ricl	/ ic

Additional interpretation of 95% CI (not needed): We are 95% confident that the (population) relative risk is between 2.70 and 4.01.

Since the 95% confidence interval does not include 1, there is sufficient evidence that the risk of diabetes differs significantly between impaired and normal glucose tolerance at baseline.

Learning Objectives

1. Understand the difference between testing for association and measuring association

- 2. Estimate the risk difference (and its confidence interval) from a contingency table and interpret the estimate.
- 3. Estimate the risk ratio (and its confidence interval) from a contingency table and interpret the estimate.

4. Estimate the odds ratio (and its confidence interval) from a contingency table and interpret the estimate.

Odds (building up to Odds Ratio)

- For a probability of success p or sometimes referred to as π), the odds of success is: $odds = \frac{p}{1-p} = \frac{\pi}{1-\pi} \qquad p = \pi$ • Example: if $\pi = 0.75$, then odds of success = 0.75• If odds > 1, it implies a success is more likely than a failure 1-0.75
 - Example: for odds = 3, we expect to observe three times as many successes as failures
- If odds is known, the probability of success can be computed

$$p = \pi = \frac{\text{odds}}{\text{odds} + 1}$$

Odds Ratio (OR)

• Odds ratio is the ratio of two odds:

$$\widehat{OR} = \frac{\widehat{Odds_1}}{\widehat{Odds_2}} = \frac{\widehat{p_1}/(1-\widehat{p_1})}{\widehat{p_2}/(1-\widehat{p_2})} = \frac{\widehat{\rho_1}(1-\widehat{\rho_2})}{\widehat{\rho_2}(1-\widehat{\rho_1})}$$

• Range: $[0,\infty]$

• Interpretation: The odds of success for "group 1" is " \widehat{OR} " times the odds of success for "group 2"

• What do values of odds ratios mean?

Odds Ratio	Clinical Meaning
$\widehat{OR} < 1$	Odds of success is smaller in group 1 than in group 2
$\widehat{OR} = 1$	Explanatory and response variables are independent
$\widehat{OR} > 1$	Odds of success is greater in group 1 than in group 2



Odds Ratio (OR)

- , clinical meaningfulners • Values of OR farther from 1.0 in a given direction represent stronger association
 - An OR = 4 is farther from independence than an OR = 2
 - An OR = 0.25 is farther from independence than an OR = 0.5
 - For OR = 4 and OR = 0.25, they are equally away from independence (because $\frac{1}{4} = 0.25$)
- We take the inverse of the OR for success of group 1 compared to group 2 to get...

p, sode

- OR for failure of group 1 compared to group 2
- OR for success of group 2 compared to group 1

OR = 1

OR = 4

Log-transformation of OR

- Like RR, sampling distribution of the odds ratio is highly skewed
 - Log transformation results in approximately normal distribution
 - Thus, compute confidence interval using normally distributed, log-transformed OR
- Approximate standard error for $\ln(\widehat{OR})$:

$$SE_{\mathrm{In}(\widehat{OR})} = \sqrt{rac{1}{n_{11}} + rac{1}{n_{12}} + rac{1}{n_{21}} + rac{1}{n_{22}}}$$

• 95% confidence interval for $\ln(\widehat{OR})$:

Explanatory	Response	Total	
Variable	Success	Failure	
1	n ₁₁	n ₁₂	n ₁
2	n ₂₁	n ₂₂	n ₂
Total	n ₊ (or n _s)	n ₋ (or n _F)	n

$${
m n}(\widehat{OR})\pm 1.96 imes SE_{{
m ln}(\widehat{OR})}$$

How do we get back to the OR scale?

• We computed confidence interval using normally distributed, log-transformed RR ($\ln(\widehat{OR})$):

$$\left(rac{\ln(\widehat{OR}) - 1.96 imes SE_{\ln(\widehat{OR})}, \ \ln(\widehat{OR}) + 1.96 imes SE_{\ln(\widehat{OR})}
ight)$$

- Now we need to exponentiate the CI to get back to interpretable values
 - Take exponential of lower and upper bounds
- 95% confidence interval for RR: two ways to display equation

$$egin{pmatrix} e^{\ln(\widehat{OR}) - 1.96 imes SE_{\ln(\widehat{OR})}}, e^{\ln(\widehat{OR}) + 1.96 imes SE_{\ln(\widehat{OR})}} \end{pmatrix} \ \left(\expig(\ln(\widehat{OR}) - 1.96 imes SE_{\ln(\widehat{OR})}ig), \ \expig(\ln(\widehat{OR}) + 1.96 imes SE_{\ln(\widehat{OR})}ig) \end{pmatrix}
ight)$$

SHS Example: Odds Ratio (1/6)

Odds ratio

Compute the point estimate and 95% confidence interval for the diabetes odds ratio between impaired and normal glucose tolerance.

Needed steps:

- 1. Compute the odds ratio
- 2. Find confidence interval of log OR
- 3. Convert back to OR
- 4. Interpret the estimate

	Diabe		
Glucose tolerance	No	Yes	Total
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

SHS Example: Odds Ratio (2/6)

Odds ratio

Compute the point estimate and 95% confidence interval for the diabetes Odds ratio between impaired and normal glucose tolerance.

	Diabe		
Glucose tolerance	No	Yes	Total
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

1. Compute the odds ratio

 $\widehat{p}_1 = 198/532 = 0.3722, \widehat{p}_2 = 128/1132 = 0.1131$

$$\widehat{OR} = rac{\widehat{p_1}/(1-\widehat{p_1})}{\widehat{p_2}/(1-\widehat{p_2})} = rac{0.3722/(1-0.3722)}{0.1131/(1-0.1131)} = 4.6499$$

SHS Example: Odds Ratio (3/6)

Odds ratio

Compute the point estimate and 95% confidence interval for the diabetes Odds ratio between impaired and normal glucose tolerance.

2. Find confidence interval of log OR

 $egin{aligned} &\ln(\widehat{OR})\pm1.96 imes SE_{\ln(\widehat{OR})} \ =&\ln(\widehat{OR})\pm z^*_{\left(1-rac{lpha}{2}
ight)} imes\sqrt{rac{1}{n_{11}}+rac{1}{n_{12}}+rac{1}{n_{21}}+rac{1}{n_{22}}} \ =&1.5368\pm1.96 imes\sqrt{rac{1}{198}+rac{1}{334}+rac{1}{128}+rac{1}{1004}} \ =&(1.2824,\ 1.7913) \end{aligned}$

	Diabe		
Glucose tolerance	No	Yes	Total
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

SHS Example: Odds Ratio (4/6)

Odds ratio

Compute the point estimate and 95% confidence interval for the diabetes Odds ratio between impaired and normal glucose tolerance.

3. Convert back to OR

 $(\exp(1.2824), \exp(1.7913))$ =(3.6053, 5.9971) DiabetesGlucose toleranceNoYesTotalImpaired334198532Normal10041281132Total13383261664

SHS Example: Odds Ratio (5/6)

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\mathbf{U}	uu	510	au	U

Compute the point estimate and 95% confidence interval for the diabetes Odds ratio between impaired and normal glucose tolerance.

1/2/3. Compute OR and 95% confidence interval



Glucose tolerance	No	Yes	Total
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

Diabetes

Pause: other option in pubh package

1 contin	ngency(case	~ glucimp, da	ata = SHS	S, digits = 3)
C	outcome			
Predictor	1 0			
Impaired	198 334			
Normal	128 1004			
				\frown
	Outcome +	Outcome -	Total	Inc risk *
Exposed +	198	334	532	37.218 (33.093 (41.482)
Exposed -	128	1004	1132	11.307 (9.521 to 13.298)
Total	326	1338	1664	19.591 (17.709 to 21.581)
Point estim Inc risk ra Inc odds ra Attrib pisk Attrib frac Attrib risk Attrib risk	ates and 95% tio tio in the expos tion in the e in the popul tion in the p	CIS: ed 1 exposed (*) ation * population (*)	6 1 4	3.291 (2.703, 4.008) 4.650 (3.605, 5.997) 25.911 (21.408, 39.113) 9.619 (63.004, 15.090) 6.284 (5.611 10.937) 2.284 (34.713, 48.976)
Uncorrected Fisher exac Wald confi CI: confid * Outcomes	l chi2 test th et test that C dence limits lence interval s per 100 popu	at OR = 1: chi DR = 1: Pr>chi2	2(1) = 15 = <0.001	4.239 Pr>chi2 = <0.001

Pearson's Chi-squared test with Yates' continuity correction

data: dat X-squared = 152.6, df = 1, p-value < 2.2e-16</pre>

SHS Example: Odds Ratio (6/6)

	Diabe		
Glucose tolerance	No	Yes	Total
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

Odds ratio

Compute the point estimate and 95% confidence interval for the diabetes Odds ratio between impaired and normal glucose tolerance.

3. Interpret the estimate



The estimated odds of diabetes for American Indians with impaired glucose tolerance at baseline is 4.65 times the odds for American Indians with normal glucose tolerance at baseline.

Additional interpretation of 95% CI (not needed): We are 95% confident that the odds ratio is between 3.61 and 6.00.

Since the 95% confidence interval does not include 1, there is sufficient evidence that the odds of diabetes differs significantly between impaired and normal glucose tolerance at baseline.

Inversing an Odds Ratio

- Some clinicians may prefer interpretations of OR > 1 instead of an OR < 1
- The transformation can easily be done by inverse
 - Remember we discussed that OR = 4 is an equivalent a strong association as OR = 0.25 (1/4)
- OR comparing group 1 to group 2 = inverse of OR comparing group 2 to group 1

$$OR_{1v2} = \frac{\hat{p}_1/(1-\hat{p}_1)}{\hat{p}_2/(1-\hat{p}_2)} = \frac{1}{\frac{\hat{p}_2/(1-\hat{p}_2)}{\hat{p}_1/(1-\hat{p}_1)}} = OR_{2v1}$$

$$OR_{2v1} = \frac{1}{OR_{1v2}}$$

Poll Everywhere Question 4

Given the estimated odds ratio (4.65) that we just calculated in our example, select the following statements that are true.

The odds of diabetes for American Indians with impaired glucose tolerance is 4.65 times the odds for 📀 American Indians with normal glucose tolerance. The odds of diabetes for American Indians with normal glucose tolerance is 0.22 times the odds for 📀 American Indians with impaired glucose tolerance The odds of diabetes for American Indians with normal glucose tolerance is 0.33 times the odds for American Indians with impaired glucose tolerance Diabetes diagnosis is less likely for those impaired glucose tolerance than those with normal glucose tolerance. Diabetes diagnosis is less likely for those normal glucose tolerance than those with impaired glucose 📀 tolerance.

45%

30%

0%

0%

24%

SHS Example: Inversing Odds Ratio

Inversing odds ratio

Compute the point estimate and 95% confidence interval for the diabetes odds ratio between **normal** and **impaired** glucose tolerance.

Needed steps:

1. Inverse point estimate and confidence interval

$$\widehat{OR} = rac{1}{4.6499} = 0.2151$$

The 95% Confidence interval is then

$$\left(rac{1}{5.9971},rac{1}{3.6053}
ight) \;=\; (0.1667, 0.2774)$$

	Diabe		
Glucose tolerance	No	Yes	Total
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

SHS Example: Inversing Odds Ratio

Inversing odds ratio

Compute the point estimate and 95% confidence interval for the diabetes odds ratio between **normal** and **impaired** glucose tolerance.

Needed steps:

1. Inverse point estimate and confidence interval

```
1 library(epitools)
2 oddsratio(x = SHS_ct, method = "wald", rev = "rows")$measure
        odds ratio with 95% C.I.
        estimate lower upper
Impaired 1.000000 NA NA
Normal 0.215059 0.1667459 0.2773702
```

	Diabe	Diabetes		
Glucose tolerance	No	Yes	Total	
Impaired	334	198	532	
Normal	1004	128	1132	
Total	1338	326	1664	

SHS Example: Inversing Odds Ratio

Inversing odds ratio

Compute the point estimate and 95% confidence interval for the diabetes odds ratio between **normal** and **impaired** glucose tolerance.

Needed steps:

2. Interpret the estimate

The estimated odds of diabetes for American Indians with normal glucose tolerance at baseline is 0.22 times the odds for American Indians with impaired glucose tolerance at baseline.

Additional interpretation of 95% CI (not needed): We are 95% confident that the odds ratio is between 0.17 and 0.28.

Since the 95% confidence interval does not include 1, there is sufficient evidence that the odds of diabetes differs significantly between impaired and normal glucose tolerance at baseline.

Glucose tolerance	No	Yes	Total
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

Learning Objectives

- 1. Understand the difference between testing for association and measuring association
- 2. Estimate the risk difference (and its confidence interval) from a contingency table and interpret the estimate.
- 3. Estimate the risk ratio (and its confidence interval) from a contingency table and interpret the estimate.
- 4. Estimate the odds ratio (and its confidence interval) from a contingency table and interpret the estimate.

pubh vs. epitools

- In pubh with contingency()
 - Get all the info at once
 - Really nice to double check how the code is interpreting your input
- In epitools with riskratio() or oddsratio()
 - Much easier to grab the numbers!
 - In Quarto you can take R code and directly put it in your text



Lesson 3: Measurement of Association for Contingency Tables