Lesson 13: Numerical Problems

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Lesson 13: Numerical Problems

Learning Objectives

- 1. Identify and troubleshoot logistic regression analysis when there are low or zero observations for the cross section of the outcome and a predictor
- 2. Identify and troubleshoot logistic regression analysis when there is complete separation between the two outcome groups
- 3. Identify and troubleshoot logistic regression analysis when there is multicollinearity between variables

Three Numerical Problems

- Issues that may cause numerical problems:
 - 1. Zero cell count
 - 2. Complete separation
 - 3. Multicollinearity

• All may cause large estimated coefficients and/or large estimated standard errors

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Zero cell count in a contingency table

• If no observations at any intersection of the covariate and outcome

• Zero cell in a contingency table should be detected in descriptive statistical analysis stage

• Example of one covariate with outcome:

Outoono	C	Tatal		
Outcome	(1)	2	3	Iotai
	(7)5	12	20	39
0	13	8	0	21
Total	20	20	20	60

Zero cell count: example (1/3)

• Example of logistic regression with **one** covariate:

1 ex1_m = glm(outcome ~ x, data = ex1, 2 family = binomial)

Outcomo	Co	Total		
Outcome	1	2	3	TOLAI
1	7	12	20	39
0	13	8	0	21
Total	20	20	20	60

Zero cell count: example (2/3)

• Example of logistic regression with **one** covariate:

1 ex1_m = glm(outcome ~ x, data = ex1, 2 family = binomial())

Coefficient estimates



Odds ratio

term	estimate	std.error	statistic	p.value	conf.low	conf.high
(Intercept)	-0.62	0.47	-1.32	0.19	-1.60	0.27
xTwo	1.02	0.65	1.57	0.12	-0.23	2.35
xThree	20.19	2,404.67	0.01	0.99	-119.00	NA
	Cex	p				

Characteristic	OR ¹	95% Cl¹	p-value
Х			
One	—	—	
Two	2.79	0.79, 10.5	0.12
Three	583,822,60	1_0.00, NA	>0.9
7 OR = Odds Ratio	o, CI = Confide	nce Interval	

Zero cell count: example (3/3)

• Example of logistic regression with **one** covariate:

1 ex1_m = glm(outcome ~ x, data = ex1, 2 family = binomial())

Coefficient estimates

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xThree	20.19	2,404.67	0.01	0.99	-119.00	NA

Coefficient estimate is large and standard error is large! Estimated odds ratio is very large and confidence interval cannot be computed.

Covariate (x) Outcome **Total** 2 3 1 7 12 39 1 20 0 13 21 8 0 Total 20 20 20 60

Odds ratio	
------------	--

Characteristic		95% Cl ¹	p-value			
х						
One	_	—				
Two	2.79	0.79, 10.5	0.12			
Three	583,822,601	0.00, NA	>0.9			
¹ OR = Odds Ratio, CI = Confidence Interval						

Ways to address zero cell

- Add one-half to each of the cell counts
 - Technically works, but not the best option
 - Rarely useful with a more complex analysis: may work for simple logistic regression
 - Nicky would say worst option because manipulating the data that does not work on individual level
 - Collapse the categories to remove the 0 cells
 - We could collapse groups 2 and 3 together if it makes clinical sense
 - Good idea if this makes clinical sense OR there is no difference between groups
 - Remove the category with 0 cells
 - This would mean we reduce the total sample size as well
 - Not a good idea: we would remove people from our dataset. Why would we do that?
 - If the variable is in ordinal scale, treat it as continuous
 - Good idea if you have seen evidence that there is a linear trend on log-odds scale

Decide on how to address zero cell (1/2)

• Look at the proportions across the predictor, X:

```
1 ggplot(data = ex1, aes(x = x, fill = outcome)) +
2 geom_bar(stat = "count", position = "fill") +
3 labs(y = "Proportion of Outcome")
```



Decide on how to address zero cell (2/2)

• Look at the proportions across the predictor, X:

```
1 ggplot(data = ex1, aes(x = x, fill = outcome)) +
2 geom_bar(stat = "count", position = "fill") +
3 labs(y = "Proportion of Outcome")
```



- Combining groups 2 and 3 together *may not be a good idea*.
- Their proportions of the outcome do not look similar.
- The predictor has an ordinal quality, so this is making me think a continuous approach might be good.

Collapse the categories of predictor

Combine groups 2 and 3:





- Based on our previous visual, I don't think this is a good idea
- Look at the estimated OR comparing group 2 to group 1 from our original model: 2.79 (95% Cl: 0.79, 10.5)
 - Looks different than the estimated OR in the above table

Remove the category with 0 cells

Remove group 3 from the data:

```
1 ex1_two = ex1 %>% filter(x != "Three")
2 ex1_two_glm = glm(outcome ~ x, data = ex1_two, family = binomial())
3 tbl_regression(ex1_two_glm, exponentiate=T) %>% as_gt() %>%
4 tab_options(table.font.size = 38)
```

Characterist	ic \mathbf{OR}^{1}	95% Cl ¹	p-value
Х			
One	_	—	
Two	2.79	0.79, 10.5	0.12
1 OR = Odds Ra	tio, CI =	Confidence	e Interval

- Not a good idea because we lose information (sample size goes down!)
- And really bad when we have other predictors!!

Treat predictor as continuous

- When we treat a predictor as continuous, we need to make sure we have **linearty between continuous predictor and log-odds**
- Cannot test this before fitting the logistic regression with the continuous predictor
 - Try taking the logit of a probability of 1... it's infinity!

```
1 ex1_cont = ex1 %>% mutate(x = as.numeric(x))
2 ex1_cont_glm = glm(outcome ~ x, data = ex1_cont, family = binomial())
3 tbl_regression(ex1_cont_glm, exponentiate=T) %>% as_gt() %>%
4 tab_options(table.font.size = 38)
```

Characteristi	$c OR^{\prime}$	95% Cl ¹	p-value
x	6.22	2.63, 18.0	<0.001
¹ OR = Odds Rat	tio, CI =	Confidence	Interval

Treat predictor as continuous: check linearity assumption

```
1 newdata = data.frame(x = c(1, 2, 3))
2 pred = predict(ex1_cont_glm, newdata, se.fit=T, type = "link")
3 LL_CI1 = pred$fit - qnorm(1-0.05/2) * pred$se.fit
4 UL_CI1 = pred$fit + qnorm(1-0.05/2) * pred$se.fit
5
6 pred_link = cbind(Pred = pred$fit, LL_CI1, UL_CI1) %>% inv.logit()
7 pred_prob = as.data.frame(pred_link) %>% mutate(x = c("One", "Two", "Three"))
```

Plotting sample and predicted probabilities



This looks pretty good. We've mostly captured the trend of the outcome proportion!

Zero cell count when we have multiple predictors

- Note that we may not see the zero count cells in a single predictor
 - But we may have issues if there is an interaction!
 - This is why I suggested we keep an eye out for cell counts below 10 in our lab!

- If you see a big <u>coefficient</u> estimate with a big standard deviation for a specific category or interaction...
 - ...this may mean that a low cell count for that category is causing you issues!



Zero cell count: summary

My suggestion is to try possible solutions in this order

1. For group with zero cell count, see if there is an adjacent group that makes sense to combine it with

2. If that does not make sense (or obscures your data) AND your data has an inherent order, then you can try treating it as continuous.

3. Remove the zero count group and all the observations in it (not a very good solution)

4. Add a half count to each cell (only works for a single predictor)

Poll Everywhere Question 1

13:23 Wed May 15	
X Join by Web PollEv.com/nickywakim275	
Can the zero cell problem apply to a continuous covariate/predicto an outcome?	r of
Yes, if both outcome groups are not observed at each continuous value	27%
Yes, we just need to use a contingency table to identify it	0%
No, there are no cells to have a zero count	55%
No, because of the linear relationship between covariate and outcome	18%
Powered by Poll Everywhere	

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Complete Separation

- Complete separation: occurs when a collection of the covariates completely separates the outcome groups
 - Example: Outcome is "gets senior discount at iHop" and the only covariate you measure is age
 - Age will completely separate the outcome
 - No overlap in distribution of covariates between two outcome groups



- Problem: the maximum likelihood estimates do not exist
 - Likelihood function is monotone
 - In order to have finite maximum likelihood estimates we must have some overlap in the distribution of the covariates in the model

Poll Everywhere Question 2



Complete Separation: example (1/3)

• We get a warning when we have complete separation





1 m1 = glm(outcome
$$\sim (x1) + x2$$
, data = ex3, family=binomial)

Warning: glm.fit: fitted probabilities numerically 0 or 1 occurred

- Outcomes of 0 and 1 are completely separated by ×
 - If $\times 2$ > 4 then outcome is 1
 - If x < 4 then outcome is 0</p>

Complete Separation: example (2/3)

Coefficient estimates:

1 tidy(m1, conf.int=T) %>% gt() %>%

```
2 tab_options(table.font.size = 35) %>%
```

3 fmt_number(decimals = 2)

term	estimate	std.error	statistic p	o.value	conf.low	conf.high
(Intercept)	-66.10	183,471.72	0.00	1.00	-10,644.72	10,512.52
x1	15.29	27,362.84	0.00	1.00	-3,122.69	NA
x2	6.24	81,543.72	0.00	1.00	-12,797.28	NA

Complete Separation: example (3/3)

Coefficient estimates:

- 1 tidy(m1, conf.int=T) %>% gt() %>%
- 2 tab_options(table.font.size = 35) %>%
- 3 fmt_number(decimals = 2)

term	estimate	std.error st	tatistic p	.value	conf.low	conf.high
(Intercept)	-66.10 1	83,471.72	0.00	1.00 -	-10,644.72	10,512.52
x1	15.29	27,362.84	0.00	1.00	-3,122.69	NA
x2	6.24	81,543.72	0.00	1.00 -	-12,797.28	NA

- Coefficient estimate of x1 is large
- Standard error of x1's coefficient is large
- But also the coefficients and standard errors for the intercept and ×2 are large!

Complete Separation

- The occurrence of complete separation in practice depends on
 - Sample size
 - Number of subjects with the outcome present
 - Number of variables included in the model

- Example: 25 observations and only 5 have "success" outcome
 - 1 variable in model may not lead to complete separation
 - More variables = more dimensions that can completely separate the observations

- In most cases, the occurrence of complete separation is not bad for clinical importance
 - But rather a numerical coincidence that causing problem for model fitting

Poll Everywhere Question 3



Complete Separation: Ways to address issue

- Collapse categorical variables in a meaningful way
 - Easiest and best if stat methods are restricted (common for collaborations)

- Exclude x1 from the model
 - Not ideal because this could lead to biased estimates for the other predicted variables in the model

- Firth logistic regression
 - Uses penalized likelihood estimation method
 - Basically takes the likelihood (that has no maximum) and adds a penalty that makes the MLE estimatable



Complete Separation: Firth logistic regression



(Intercept)	2
x1	2
x2	2

Method: 1-Wald, 2-Profile penalized log-likelihood, 3-None

```
Likelihood ratio test=5.505687 on 2 df, p=0.06374636, n=8
Wald test = 3.624899 on 2 df, p = 0.1632538
```



p

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Multicollinearity

• **Multicollinearity** happens when one or more of the covariates in a model can be predicted from other covariates in the same model

• This will cause unreliable coefficient estimates for some covariates in logistic regression, as in an ordinary linear regression

- Looking at correlations among pairs of variables is helpful but not enough to identify multicollinearity problem
 - Because multicollinearity problems may involve relationships among more than two covariates

Multicollinearity: example (1/4)

• Table below is a simulated data with

•
$$x_1 \sim \text{Normal}(0, 1)$$

• $x_2 = x_1 + \text{Uniform}(0, 0.1)$
• $x_2 = 1 + \text{Uniform}(0, 0.01)$

• Therefore, x_1 and x_2 are highly correlated, and x_3 is nearly collinear with the constant term

Independent Variables and Constant							
Subject	<i>x</i> ₁	<i>x</i> ₂	<i>x</i> ₃	у			
1	0.225	0.231	1.026	0			
2	0.487	0.489	1.022	1			
3	-1.080	-1.070	1.074	0			
4	-0.870	-0.870	1.091	0			
5	-0.580	-0.570	1.095	0			
6	-0.640	-0.640	1.010	0			
7	1.614	1.619	1.087	0			
8	0.352	0.355	1.095	1			
9	-1.025	-1.018	1.008	0			
10	0.929	0.937	1.057	1			

Table 4.38Data Displaying Near Collinearity Among theIndependent Variables and Constant

Multicollinearity: example (2/4)

- Four logistic regression models using data in the previous slide
- Consequence of multicollinearity: large coefficient estimates and/or standard errors

Regression Models to the Data in Table 4.38								
Var.	Coeff.	Std. Err.	Coeff.	Std. Err.	Coeff.	Std. Err.	Coeff.	Std. Err.
$\frac{x_1}{x_2}$	1.4	1.0	104.2 -103.4	256.2 256.0			79.8 -78.3	272.6 272.5
x_3 Cons.	-1.0	0.8	-0.3	1.3	1.8 - 2.7	20.0 21.1	-11.1 11.4	206.6 27.8
	M (\$)\$ (\$)	1: xil omly	Model 2	: x <u>1</u> and x2	Model -	3:53 only	[23(<u>636633</u>)	석유) 제비가 10%) 기억화

. .

Multicollinearity: example (3/4)

Table 4.39

• Four logistic regression models using data in the previous slide

Regression Models to the Data in Table 4.38

• Consequence of multicollinearity: large coefficient estimates and/or standard errors

			Lange lociofiliteiren (la sinul cizi		La Mara (2 E		ar sin	arge deniciand
	Model 1: x1 only		Model 2: x1 and x2		Model 3: x3 only		Mode	el 4: x1, x2, x3
x_3 Cons.	-1.0	0.8	-0.3	1.3	$\begin{pmatrix} 1.8 \\ -2.7 \end{pmatrix}$	$\left(\begin{array}{c} 20.0\\21.1\end{array}\right)$	-11.1 11.4	206.6 27.8
$\begin{array}{c} x_1 \\ x_2 \end{array}$	1.4	1.0	104.2 -103.4	256.2 256.0		\frown	79.8 -78.3	272.6 272.5
Var.	Coeff.	Std. Err.	Coeff.	Std. Err.	Coeff.	Std. Err.	Coeff.	Std. Err.
	_		_				_	

Estimated Coefficients and Standard Errors from Fitting Logistic

Multicollinearity: example (4/4)

- Four logistic regression models using data in the previous slide
- Consequence of multicollinearity: large coefficient estimates and/or standard errors

$\begin{array}{c ccccccccccccccccccccccccccccccccccc$		Large coefficients and SE
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$		nly Model 4: x1, x2, x3
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	s ons.	$\begin{array}{cccc} 0 & -11.1 & 206.6 \\ 1 & 11.4 & 27.8 \end{array}$
	2	79.8272.6-78.3272.5
Var. Coeff. Std. Err. Coeff. Std. Err. Coeff. Std. Err. Coeff. Std. Err.	ar.	Err. Coeff. Std. Err.

Table 4.39Estimated Coefficients and Standard Errors from Fitting LogisticRegression Models to the Data in Table 4.38

Multicollinearity: how to detect

- Multicollinearity only involves the covariates
 - No specific issues to logistic regression (vs. linear regression)
 - Techniques from 512/612 work well for logistic regression model

• In more complicated dataset/analysis, we may not be able to detect multicollinearity using the coefficient estimates/SE

• Variance inflation factor (VIF) approach: well-known approach to detect multicollinearity

Variance Inflation Factor (VIF) Approach

- Computed by regressing each variable on all the other explanatory variables
 - For example: $E(x_1)x_2, x_3, \ldots) = \alpha_0 + \alpha_1x_2 + \alpha_2x_3$
- Calculate the coefficient of determination, R^2
 - Proportion of the variation in x_1 that is predicted from $x_2, x_3,...$

$$VIF = \frac{1}{1 \, \mathbf{G} \, R^2}$$

- Each covariate has its own VIF computed
- Get worried for multicollinearity if VIF > 10
- Sometimes VIF approach may miss serious multicollinearity
 - Same multicollinearity we wish to detect using VIF can cause numerical problems in reliably estimating R^2

sinteractions or > correlated Xi Multicollinearity: Ways to address the issue

- Exclude the redundant variable from the model
- Scaling and centering variables
 - When you have transformed a continuous variable
- Other modeling approach (outside scope of this class)
 - Ridge regression
 - Principle component analysis

Please take a look at the BSTA 512/612 lesson that included multicollinearity

Poll Everywhere Question 4



Remore X3 Remore either X1 or X2

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